



## Probability Theory

## The survival probability of a critical branching process in a Markovian random environment

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## ABSTRACT

In this Note, we first prove a local limit theorem for a semi-Markov chain and then apply it to study the asymptotic behavior of the survival probability of a critical branching process in Markovian random environment.

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## R É S U M É

Dans cette Note, nous montrons d'abord un théorème de la limite locale pour une chaîne semi-Markovienne. Nous appliquons ensuite ce résultat pour étudier le comportement asymptotique de la probabilité de survie d'un processus de branchement critique dans un milieu aléatoire Markovien.

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## 1. Introduction and main results

The study of branching processes in Markovian random environment has been developed by several authors, in particular by K.B. Athreya and S. Karlin [1]. However, the asymptotic behavior of the survival probability of such a process is not yet known. In this Note we handle this problem in the case of a critical branching process.

Consider the following model:  $X = (X_n)_{n \geq 0}$  is an irreducible and aperiodic Markov chain on a finite space  $E$  with transition matrix  $P$ . The chain  $X$  has a unique invariant probability  $\nu$ . We denote by  $G$  the set of generating functions of probability measures on  $\mathbb{N}$ , equipped with the topology of simple convergence on  $[0, 1]$ .  $\mathcal{B}(G)$  is the Borel  $\sigma$ -algebra on  $G$ . In addition, we define a Markov chain  $(M_n)_{n \geq 0} = (g_n, X_n)_{n \geq 0}$  with values in  $G \times E$  and with transition probability  $Q$  defined by

$$Q\{(g, i), (A \times \{j\})\} = P(i, j)\bar{F}(i, j, A), \quad \text{for } (g, i) \in G \times E, A \in \mathcal{B}(G),$$

where  $\bar{F}$  is a transition probability from  $E \times E$  in the set of probabilities on  $G$ . The Markov chain  $(M_n)_{n \geq 0}$  is called the *environment process*. Let  $\Omega = (G \times E)^{\mathbb{N}}$  and  $\mathcal{F} = \bigotimes_{\mathbb{N}} (\mathcal{B}(G) \otimes \mathcal{P}(E))$ . We denote by  $\mathbb{P}_{(g,i)}$  the unique probability on  $(\Omega, \mathcal{F})$ , such that for any  $(g, i) \in G \times E$ , any  $n \geq 1$  and any bounded measurable function  $f : (G \times E)^n \rightarrow \mathbb{R}$ , we get

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$$\int_{\Omega} f(M_0(\omega), M_1(\omega), \dots, M_n(\omega)) \mathbb{P}_{(g,i)}(d\omega) = \sum_{(j_1, j_2, \dots, j_n) \in E^n} P(i, j_1) \cdots P(j_{n-1}, j_n) \int_{G^n} f((g, i), (g_1, j_1), \dots, (g_n, j_n)) \bar{F}(i, j_1, dg_1) \cdots \bar{F}(j_{n-1}, j_n, dg_n).$$

To simplify the notations,  $\mathbb{P}_{(id,i)}$  will be denoted by  $\mathbb{P}_i$  and  $\mathbb{E}_i$  is its corresponding expectation.

Given  $(M_n)_{n \geq 0}$ , we define now the branching process  $(Z_n)_{n \geq 0}$  such that  $Z_0 = 1$  and the generating function of  $Z_n$  is

$$g_0 \circ g_1 \circ \cdots \circ g_{n-1}(s) := G_n(s), \quad 0 \leq s < 1.$$

Therefore, given  $(M_n)_{n \geq 0}$ , the survival probability of the branching process  $(Z_n)_{n \geq 0}$  at time  $n$  is

$$1 - G_n(0) := q_n.$$

Due to the Markov property of the probability  $\mathbb{P}_i$ , we have for  $(i, j) \in E \times E$  and  $n \geq 1$ ,

$$\mathbb{P}_i(Z_n > 0, X_n = j) = \mathbb{E}_i(q_n P(X_{n-1}, j)).$$

Let's consider  $h : G \rightarrow \bar{\mathbb{R}}_+$ ,  $g \mapsto h(g) := g'(1)$ . The image of the probability  $\bar{F}(i, j, dx)$  by the map  $h$  is denoted by  $F(i, j, dx)$ . We assume in this paper the following hypotheses (H):

(H1) there exist  $\alpha > 0$ , such that for all  $\lambda \in \mathbb{C}$  satisfying  $|\operatorname{Re} \lambda| \leq \alpha$ , we have

$$\sup_{(i,j) \in E \times E} |\hat{F}(i, j, \lambda)| < +\infty, \quad \text{where } \hat{F}(i, j, \lambda) = \int_{\mathbb{R}} e^{\lambda t} F(i, j, dt);$$

(H2) there exist  $n_1 \geq 1$  and  $(i_0, j_0) \in E \times E$ , such that the measure  $\mathbb{P}_{i_0}(X_{n_1} = j, S_{n_1} \in dx)$  has an absolutely continuous component with respect to the Lebesgue measure on  $\mathbb{R}$ ;

(H3)  $\sum_{(i,j) \in E \times E} \nu(i) P(i, j) \int_{\mathbb{R}} t F(i, j, dt) = 0$ .

By [1], the hypothesis (H3) implies

$$\mathbb{P}_\nu(Z_n = 0) \rightarrow 1, \quad \text{as } n \rightarrow +\infty.$$

Such a branching process  $(Z_n)_{n \geq 0}$  is called *critical*.

Let us introduce some notations: set

$$\eta_{k,n} := f_k(g_{k+1,n}(0)),$$

where

$$f_k(s) := \frac{1}{1 - g_k(s)} - \frac{1}{g'_k(1)(1 - s)}, \quad \text{for } 0 \leq s < 1 \quad \text{and} \quad g_{k,n} := g_k \circ g_{k+1} \circ \cdots \circ g_{n-1}, \quad \text{for } 0 \leq k \leq n - 1;$$

$$S_n := Y_0 + Y_1 + \cdots + Y_{n-1}, \quad \text{for } n \geq 1, \quad \text{with } S_0 := 0 \quad \text{and} \quad Y_k := \ln g'_k(1), \quad \text{for } k \geq 0.$$

Then we can obtain the following formula [4]:

$$q_n^{-1} = \exp(-S_n) + \sum_{k=0}^{n-1} \eta_{k,n} \exp(-S_k). \tag{1}$$

**Theorem 1.1.** Under hypotheses (H), for any  $(i, j) \in E \times E$ , there exists a constant  $\beta_{i,j} > 0$ , such that

$$\lim_{n \rightarrow +\infty} \sqrt{n} \mathbb{P}_i(Z_n > 0, X_n = j) = \beta_{i,j}. \tag{2}$$

For  $n \geq 0$ , we set

$$m_n = \min\{S_0, S_1, \dots, S_n\}.$$

The proof of Theorem 1.1 is based on the following local limit theorem:

**Theorem 1.2.** Under hypotheses (H), for any  $(i, j) \in E \times E$  and  $x \geq 0$ , we get

$$\lim_{n \rightarrow +\infty} \sqrt{n} \mathbb{P}_i(m_n \geq -x, X_n = j) = h_{i,j}(x) > 0, \tag{3}$$

where  $h_{i,j}$  is an increasing harmonic function for  $(S_n, X_n)_{n \geq 0}$  on  $\mathbb{R}_+ \times E$ .

Furthermore, there exists a constant  $\sigma^2 > 0$ , such that

$$h_{i,j}(x) \sim \sqrt{\frac{2}{\sigma^2}} \nu(j)x, \quad x \rightarrow +\infty. \tag{4}$$

J. Geiger and G. Kersting [4], Y. Guivarc’h, E. Le Page and Q. Liu [5] proved an analog of Theorem 1.1 in the case of i.i.d. environment under weaker moment assumptions and without any hypotheses of absolute continuity of type (H2).

In the case when  $E$  contains one single point, Theorem 1.2 extends the local limit theorem for the minimum of a random walk on  $\mathbb{R}$  (see also [2]). Theorem 1.2 improves the results of E.L. Presman [6], especially we prove that for any  $(i, j) \in E \times E$ , the limit function  $h_{i,j}$  defined in (3) does not vanish and we specify its asymptotic behavior as  $x \rightarrow +\infty$ .

**2. Sketch of proofs**

To prove Theorem 1.2, we make use of a factorization method due to E.L. Presman [6]. We denote by  $L_\infty(E)$  the space of bounded function on  $E$ , equipped with the uniform norm. We define matrices  $\mathcal{P}B_z(\lambda)$ ,  $\mathcal{Q}C_z(\lambda)$ , and Fourier–Laplace operators  $P(\lambda)$  on  $L_\infty(E)$  as follows:

$$\begin{aligned} \mathcal{P}B_z(\lambda) &= \left[ \sum_{n=1}^{+\infty} z^n \mathbb{E}_i(e^{\lambda S_n}; S_1 > S_n, S_2 > S_n, \dots, S_{n-1} > S_n, S_n < 0; X_n = j) \right]_{i,j}, \\ \mathcal{Q}C_z(\lambda) &= \left[ \sum_{n=1}^{+\infty} z^n \mathbb{E}_i(e^{\lambda S_n}; S_1 \geq 0, S_2 \geq 0, \dots, S_{n-1} \geq 0, S_n \geq 0; X_n = j) \right]_{i,j}, \\ P(\lambda)\varphi(i) &= \sum_{j \in E} P(i, j)\varphi(j)\widehat{F}(i, j, \lambda) = \sum_{j \in E} P(i, j)\varphi(j) \int_{\mathbb{R}} e^{\lambda t} F(i, j, dt), \end{aligned}$$

for  $\varphi \in L_\infty(E)$ ,  $i \in E$  and  $|\operatorname{Re} \lambda| < \alpha$ .

We first prove that the matrix  $H(z, \lambda) := \sqrt{1-z} [\sum_{n=0}^{+\infty} z^n \mathbb{E}_i(e^{\lambda m_n}; X_n = j)]$  can be factorized as follows: for  $|z| < 1$ ,  $\operatorname{Re} \lambda = 0$ ,

$$H(z, \lambda) = [I + \mathcal{P}B_z(\lambda)] \sqrt{1-z} [I + \mathcal{Q}C_z(0)].$$

In addition, we have the following identity, which is analogous to the well-known Wiener–Hopf factorization [3],

$$(I - zP(\lambda))^{-1} = [I + \mathcal{P}B_z(\lambda)][I + \mathcal{Q}C_z(\lambda)], \quad |z| < 1, \operatorname{Re} \lambda = 0. \tag{5}$$

Then using E.L. Presman’s factorization theory [6] and especially analytical properties of such factorization, we can prove that

- (1) for  $\operatorname{Re} \lambda > 0$ , the function  $[I + \mathcal{P}B_z(\lambda)]$  is analytic with respect to  $z$  in  $D_{\rho,\theta} = \{z; z \neq 1, |\arg(z-1)| \geq \theta > 0, |z| < \rho\}$ ,  $\rho > 1, 0 < \theta < \pi/2$  and admits an analytical extension to the boundary of  $D_{\rho,\theta}$ ;
- (2) as  $\lambda \rightarrow 0$ , the limit of  $\lambda[I + \mathcal{P}B_1(\lambda)]$  exists;
- (3)  $[I + \mathcal{Q}C_z(0)]$  is analytic with respect to  $z$  in  $D_{\rho,\theta}$ . Furthermore,  $\sqrt{1-z}[I + \mathcal{Q}C_z(0)]$  is bounded in  $D_{\rho,\theta}$  and admits a limit as  $z \rightarrow 1$ .

So for any  $(i, j) \in E \times E$ , the limit of  $[H(z, \lambda)]_{i,j}$  as  $z \rightarrow 1$  exists and is denoted by  $[H(\lambda)]_{i,j}$ , from which, leads to (3), using complex analysis argument. Moreover, we get

$$\lim_{\lambda \rightarrow 0^+} \lambda [H(\lambda)]_{i,j} = \sqrt{\frac{2}{\sigma^2}} \nu(j) > 0,$$

where  $\sigma^2$  is a positive constant. Using Tauberian theorem [3], we obtain (4).

The proof of Theorem 1.1 is similar to the one in [4]. An important step is to check, using Theorem 1.2 and the formula (1), that for any  $\varrho > 1, x \geq 0$  and  $i \in E$ , we have

$$\limsup_{m \rightarrow +\infty} \limsup_{n \rightarrow +\infty} \sqrt{n} \mathbb{P}_i(Z_m > 0, Z_n = 0, m_{\varrho n} \geq -x) = 0. \tag{6}$$

### 3. More information on the proof of Theorem 1.1

For every  $x \geq 0$  and every  $(i, j) \in E \times E$ , let denote by  $\widehat{\mathbb{E}}_{(i,j,x)}$  the expectation corresponding to the unique probability  $\widehat{\mathbb{P}}_{(i,j,x)}$  on  $(\Omega, \mathcal{F})$  such that for every integer  $n \geq 1$  and every measurable, bounded function  $f : (G \times E)^n \rightarrow \mathbb{R}$ , we have

$$\int_{\Omega} f(M_1(\omega), \dots, M_n(\omega)) \widehat{\mathbb{P}}_{(i,j,x)}(d\omega) = \frac{1}{h_{i,j}(x)} \int_{\Omega} f(M_1(\omega), \dots, M_n(\omega)) h_{X_n(\omega), j}(x + S_n(\omega)) \mathbb{P}_i(d\omega).$$

Set  $q_{\infty} = \lim_{n \rightarrow +\infty} q_n$ . Using the equality (6) and Theorem 1.2, we can establish that

$$\lim_{n \rightarrow +\infty} \sqrt{n} \mathbb{P}_i(Z_n > 0, X_n = j) = \lim_{x \rightarrow +\infty} h_{i,j}(x) \widehat{\mathbb{E}}_{(i,j,x)}(q_{\infty}).$$

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